

Expectation of joint distribution

Let (X_1, X_2) be a random vector and let $Y = g(X_1, X_2)$ for some real-valued function g .

Suppose (X_1, X_2) is of continuous type.

The $E(Y)$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1, x_2) \cdot f_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

if $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |g(x_1, x_2)| f_{X_1, X_2}(x_1, x_2) dx_1 dx_2 < \infty$

In case, (X_1, X_2) is discrete,

$$E(Y) = \sum_{x_1} \sum_{x_2} g(x_1, x_2) P_{X_1, X_2}(x_1, x_2)$$

if $\sum_{x_1} \sum_{x_2} |g(x_1, x_2)| P_{X_1, X_2}(x_1, x_2) < \infty$

Now, we shall show that E is a linear operator.

Theorem. Let $Y_1 = g_1(X_1, X_2)$ and $Y_2 = g_2(X_1, X_2)$ be random variables such that $E(Y_1)$ and $E(Y_2)$ exist.

Then, for all real numbers k_1 & k_2 ,

$$E(k_1 Y_1 + k_2 Y_2) = k_1 E(Y_1) + k_2 E(Y_2).$$

Proof: We shall first show that $E(k_1 Y_1 + k_2 Y_2)$ exist. We assume that (X_1, X_2) is of continuous type.

Consider $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |k_1 g_1(x_1, x_2) + k_2 g_2(x_1, x_2)| f_{X_1, X_2}(x_1, x_2) dx_1 dx_2$

$$\leq |k_1| \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |g_1(x_1, x_2)| f_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

$$+ |k_2| \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} |g_2(x_1, x_2)| f_{X_1, X_2}(x_1, x_2) dx_1 dx_2 \quad \dots (i)$$

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using the triangle inequality and linearity of integrals.

Since $E(Y_1)$ and $E(Y_2)$ exist,

both the integrals on the R.H.S. of (i) are finite ($< \infty$), and hence, the the integral on the L.H.S. of (i) is $< \infty$.

$\Rightarrow E(k_1 Y_1 + k_2 Y_2)$ exist.

By Definition,

$$E(k_1 Y_1 + k_2 Y_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (k_1 \delta_1(x_1, x_2) + k_2 \delta_2(x_1, x_2)) \times f_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

$$= k_1 \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \delta_1(x_1, x_2) f_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

$$+ k_2 \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \delta_2(x_1, x_2) f_{X_1, X_2}(x_1, x_2) dx_1 dx_2$$

$$= k_1 E(Y_1) + k_2 E(Y_2) \quad (\text{by linearity of integrals})$$

i.e. the desired result. //

Remark. In case, we have a function $g(X_1)$ of X_1 (or $g(X_2)$ of X_2), the expected value of it can be found in two ways:

$$E(g(X_1)) = \begin{cases} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x_1) f_{X_1, X_2}(x_1, x_2) dx_1 dx_2, & (\text{by defn.}) \\ \int_{-\infty}^{\infty} g(x_1) f_{X_1}(x_1) dx_1 & \leftarrow \text{marginal pdf of } X_1. \end{cases}$$

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Consider the example:

Example 1: Let X_1 and X_2 have the pdf

$$f(x_1, x_2) = \begin{cases} 8x_1x_2 & 0 \leq x_1 \leq x_2 \leq 1 \\ 0 & \text{elsewhere} \end{cases}$$

Evaluate $E(7X_1X_2^2 + 5X_2)$.

Solution: We know that E is a linear operator, hence

$$E(7X_1X_2^2 + 5X_2) = 7 \cdot E(X_1X_2^2) + 5E(X_2).$$

Now, by definition,

$$E(X_1X_2^2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_1x_2^2 f(x_1, x_2) dx_1 dx_2$$

$$= \int_0^1 \int_0^{x_2} 8x_1x_2^3 dx_1 dx_2$$

$$= \int_0^1 \frac{8}{2} x_2^6 dx_2 = \frac{8}{26} \quad \dots (i)$$

and $E(X_2) = \int_0^1 \left(\int_0^{x_2} x_2 (8x_1x_2) dx_1 \right) dx_2$ (by defn)

$$= \int_0^1 4x_2^4 dx_2 = \frac{4}{5} \quad \dots (ii)$$

Thus, $E(7X_1X_2^2 + 5X_2) = 7 \times \frac{8}{26} + 5 \times \frac{4}{5} = \frac{20}{13} //$

Note: We observe that the marginal pdf of X_2 ,

$$f_2(x_2) = \int_0^{x_2} f(x_1, x_2) dx_1$$

$$= \begin{cases} \int_0^{x_2} 8x_1x_2 dx_1 & \text{if } 0 \leq x_2 \leq 1 \\ 0 & \text{else} \end{cases}$$

$$= \begin{cases} 4x_2^2 & \text{if } 0 \leq x_2 \leq 1 \\ 0 & \text{else} \end{cases}$$

$\therefore E(X_2) = \int_0^1 x_2 f_2(x_2) dx_2 = \int_0^1 4x_2^4 dx_2 = \frac{4}{5}$, as (iii) //

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Example 2: Let X_1, X_2 be two random variables with the joint pmf, $p(x_1, x_2) = \begin{cases} (x_1 + x_2)/12 & \text{for } x_1 = 1, 2, \\ & x_2 = 1, 2 \\ 0 & \text{elsewhere} \end{cases}$

(i) Compute $E(X_1)$, $E(X_1^2)$, $E(X_2)$, $E(X_2^2)$, and $E(X_1 X_2)$.

(ii) Is $E(X_1 X_2) = E(X_1)E(X_2)$?

(iii) Find $E(2X_1 - 6X_2^2 + 7X_1 X_2)$.

Solution: (i) $E(X_1) = \sum_{x_1=1}^2 \sum_{x_2=1}^2 x_1 p(x_1, x_2)$

$$= \sum_{x_1=1}^2 \sum_{x_2=1}^2 \frac{x_1 (x_1 + x_2)}{12}$$

$$= \sum_{x_1=1}^2 \left[\frac{x_1(x_1+1)}{12} + \frac{x_1(x_1+2)}{12} \right]$$

$$= \left[\frac{1 \cdot 2}{12} + \frac{1 \cdot 3}{12} \right] + \left[\frac{2 \cdot 3}{12} + \frac{2 \cdot 4}{12} \right]$$

$$= \frac{5}{12} + \frac{14}{12} = \frac{19}{12}$$

$$E(X_1^2) = \sum_{x_1=1}^2 \sum_{x_2=1}^2 x_1^2 p(x_1, x_2)$$

$$= \sum_{x_1=1}^2 \sum_{x_2=1}^2 \frac{x_1^2 (x_1 + x_2)}{12}$$

$$= \sum_{x_1=1}^2 \left[\frac{x_1^3 + x_1^2}{12} + \frac{x_1^3 + 2x_1^2}{12} \right]$$

$$= \left[\frac{1^3 + 1^2}{12} + \frac{1^3 + 2 \cdot 1^2}{12} \right] + \left[\frac{2^3 + 2^2}{12} + \frac{2^3 + 2 \cdot 2^2}{12} \right]$$

$$= \left[\frac{2}{12} + \frac{3}{12} \right] + \left[\frac{12}{12} + \frac{16}{12} \right]$$

$$= \frac{5}{12} + \frac{28}{12} = \frac{33}{12} = \frac{11}{4}$$

Similarly, $E(X_2) = \sum_{x_2=1}^2 \sum_{x_1=1}^2 \frac{x_2 (x_1 + x_2)}{12} = \frac{19}{12}$, and

$$E(X_2^2) = \frac{11}{4}$$

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$$\begin{aligned}
 \text{and } E(X_1 X_2) &= \sum_{x_1=1}^2 \sum_{x_2=1}^2 x_1 x_2 \frac{(x_1 + x_2)}{12} \\
 &= \sum_{x_1=1}^2 \left[\frac{x_1(x_1+1)}{12} + \frac{2x_1(x_1+2)}{12} \right] \\
 &= \left[\frac{2}{12} + \frac{2 \cdot 3}{12} \right] + \left[\frac{2 \cdot 3}{12} + \frac{4 \cdot 4}{12} \right] \\
 &= \frac{8}{12} + \frac{22}{12} = \frac{30}{12} = \frac{5}{2}
 \end{aligned}$$

(ii) we have

$$E(X_1) E(X_2) = \frac{19}{12} \times \frac{15}{12} = \frac{361}{144} \neq \frac{5}{2} = E(X_1 X_2)$$

(iii) $E(2X_1 - 6X_2^2 + 7X_1 X_2)$

$$= 2E(X_1) - 6E(X_2^2) + 7E(X_1 X_2), \text{ as } E \text{ is a linear operator}$$

$$= 2 \times \frac{19}{12} - 6 \times \frac{11}{4} + 7 \times \frac{5}{2}$$

$$= \frac{19}{6} - \frac{33}{2} + \frac{35}{2} = \frac{19 - 99 + 105}{6} = \frac{25}{6}$$

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Moment Generating Function (MGF) of a Random Vector:

We define the MGF of a random vector $X = (X_1, X_2)$

$$M_{X_1, X_2}(t_1, t_2) \text{ by } M_{X_1, X_2}(t_1, t_2) = E(e^{t_1 X_1 + t_2 X_2})$$

if $E(e^{t_1 X_1 + t_2 X_2})$ exists for $|t_1| < h_1$ and $|t_2| < h_2$,

where h_1 and h_2 are positive.

if $t = (t_1, t_2)'$, then we can write the MGF of X

$$M_{X_1, X_2}(t) = E[e^{t'X}]$$

Note that, the mgfs of X_1 and X_2 are

$$M_{X_1, X_2}(t_1, 0) \text{ and } M_{X_1, X_2}(0, t_2), \text{ respectively.}$$

Example 3. Let the continuous-type random variables X and Y have the joint pdf

$$f(x, y) = \begin{cases} e^{-x-y}, & 0 < x < \infty, 0 < y < \infty \\ 0, & \text{elsewhere.} \end{cases}$$

Find the joint mgf of X and Y . Is $M_{X_1, X_2}(t_1, t_2) = M(t_1, 0) M(0, t_2)$?

Solution: By definition,

$$M_{X, Y}(t_1, t_2) = E(e^{t_1 X + t_2 Y})$$

$$= \int_0^{\infty} \int_0^{\infty} e^{t_1 x + t_2 y} f(x, y) dx dy$$

$$= \int_0^{\infty} \int_0^{\infty} e^{t_1 x + t_2 y} e^{-x-y} dx dy$$

$$= \int_0^{\infty} e^{(t_2-1)y} \left(\int_0^{\infty} e^{(t_1-1)x} dx \right) dy$$

$$= \int_0^{\infty} e^{(t_2-1)y} \times \frac{1}{(1-t_1)} dy \quad \text{if } t_1 < 1$$

$$= \frac{1}{(1-t_1)(1-t_2)} \quad \text{if } t_1 < 1 \text{ and } t_2 < 1$$

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Next, the mgfs of the marginal distributions of X and Y are

$$M(t_1, 0) = E(e^{t_1 X})$$

$$= \int_0^{\infty} \int_0^{\infty} e^{t_1 x} \cdot e^{-x-y} dx dy$$

$$= \int_0^{\infty} e^{-y} \left(\int_0^{\infty} e^{-(1-t_1)x} dx \right) dy$$

$$= \int_0^{\infty} e^{-y} \cdot \left(\frac{1}{1-t_1} \right) dy \quad \text{if } t_1 < 1$$

$$= \frac{1}{(1-t_1)} \quad \text{if } t_1 < 1$$

and $M(0, t_2) = E(e^{t_2 Y})$

$$= \int_0^{\infty} \int_0^{\infty} e^{(t_2-1)y} e^{-x} dy dx$$

$$= \frac{1}{(1-t_2)} \quad \text{if } t_2 < 1$$

we have $M(t_1, 0) M(0, t_2)$

$$= \frac{1}{(1-t_1)} \cdot \frac{1}{(1-t_2)} \quad \text{if } t_1 < 1 \text{ and } t_2 < 1.$$

$$= M_{X,Y}(t_1, t_2),$$

i.e. the result //

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Example 4. Let X_1, X_2 be two random variables with joint pmf

$$p(x_1, x_2) = \begin{cases} \left(\frac{1}{2}\right)^{x_1+x_2} & \text{for } 1 \leq x_i < \infty, i=1,2 \\ 0 & \text{elsewhere} \end{cases}$$

where x_1 and x_2 are integers.

- (i) Determine the joint mgf of X_1, X_2 .
 (ii) Show that $M(t_1, t_2) = M(t_1, 0)M(0, t_2)$.

Solution:

$$\begin{aligned} M_{X_1, X_2}(t_1, t_2) &= E\left(e^{t_1 X_1 + t_2 X_2}\right) \\ &= \sum_{x_1=1}^{\infty} \sum_{x_2=1}^{\infty} e^{t_1 x_1 + t_2 x_2} \left(\frac{1}{2}\right)^{x_1+x_2} \\ &= \sum_{x_1=1}^{\infty} \left(e^{t_1} \frac{1}{2}\right)^{x_1} \sum_{x_2=1}^{\infty} \left(e^{t_2} \frac{1}{2}\right)^{x_2} \\ &= \left[\frac{1}{1 - \frac{1}{2} e^{t_1}} - 1 \right] \left[\frac{1}{1 - \frac{1}{2} e^{t_2}} - 1 \right] \end{aligned}$$

provided $t_i < \log 2, i=1, 2$.

$$\text{Next, } M(t_1, 0) = \sum_{x_1=1}^{\infty} \sum_{x_2=1}^{\infty} e^{t_1 x_1} \left(\frac{1}{2}\right)^{x_1} \left(\frac{1}{2}\right)^{x_2}$$

$$= \left(\sum_{x_1=1}^{\infty} \left(e^{t_1} \frac{1}{2}\right)^{x_1} \right) \left(\sum_{x_2=1}^{\infty} \left(\frac{1}{2}\right)^{x_2} \right)$$

$$= \left[\frac{1}{1 - \frac{1}{2} e^{t_1}} - 1 \right] \times \left(\frac{1}{1 - \frac{1}{2}}\right)$$

$$= \left[\frac{1}{1 - \frac{1}{2} e^{t_1}} - 1 \right], \text{ and } (t_1 < \log 2)$$

$$\text{Similarly, } M(0, t_2) = \left[\frac{1}{1 - \frac{1}{2} e^{t_2}} - 1 \right], \text{ provided } t_2 < \log 2$$

Hence, $M(t_1, t_2) = M(t_1, 0)M(0, t_2)$, i.e. the result //

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